

SETUP

- Let \mathbb{A} be a finite set and consider a subset \mathcal{P} of $M_+^1(\mathbb{A})$.
- Generally, we refer to \mathcal{P} as our **model**, and we wish to determine a probability distribution P^* on \mathbb{A} such that

$$\sup_{P \in \mathcal{P}} D(P \| P^*) = \inf_{Q \in M_+^1(\mathbb{A})} \sup_{P \in \mathcal{P}} D(P \| Q)$$

where D is the Kullback-Leibler divergence. When such a distribution exists, we say that P^* is the **minimax predictor** for the model \mathcal{P} .

The goal is to determine, in closed form, such a minimax predictor for as many explicit (families of) models as possible.

ORDER MODELS

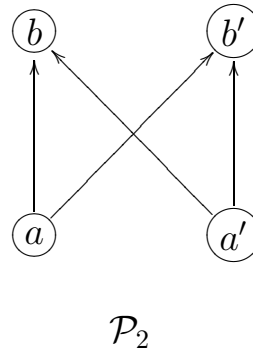
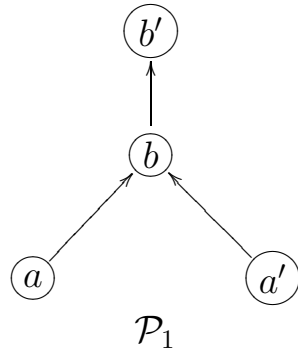
- Of specific interest are **order consistent models**,. Consider a (usually partial) ordering ' \leq ' on \mathbb{A} and let \mathcal{P} be the model of probability distributions on \mathbb{A} satisfying

$$a \leq b \Rightarrow P(a) \geq P(b)$$

- Pictorially we depict an ordering on \mathbb{A} by a graph where $\textcircled{a} \longrightarrow \textcircled{b}$ means that $a \leq b$, and we leave out arrows that are redundant due to transitivity.
- It can be proved that **for order models, the minimax predictor always exists, and it is unique.**

SIMPLE EXAMPLES

- Consider the alphabet $\mathbb{A} = \{a, a', b, b'\}$. On this we may consider, among others, 2 order models given by the graphs below.



- Remarkably, it turns out that the **minimax predictors are equal for the two models**. Specifically, in both cases, $P^*(a) = P^*(a') = \frac{16}{34}$ and $P^*(b) = P^*(b') = \frac{1}{34}$.

This is not so hard to see, once we have noticed that for \mathcal{P}_1 , the minimax predictor assigns equal probability to b and b' . (to see this, of course we need to actually calculate it, but this is doable!)

BASIC TECHNIQUES

- Our basic strategy for determining minimax predictors for order models is to make (sometimes educated) guesses and then apply the following
- **Kuhn-Tucker criterion:** If P^* is a (finite) convex mixture $P^* = \sum \alpha_k P_k$ of order consistent distributions, and

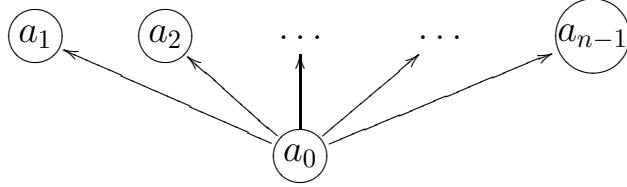
$$\begin{aligned} D(P_k \| P^*) &= R \quad \text{for all } k, \\ D(P \| P^*) &\leq R \quad \text{for all } P \in \mathcal{P}, \end{aligned}$$

then P^* is the minimax predictor for \mathcal{P} .

- The uniqueness result mentioned above is in fact even more powerful, as it says that there always exists some set of order consistent distributions/weights yielding a minimax predictor through the **Kuhn-Tucker criterion**.
- A set $\Delta \subseteq \mathbb{A}$ is said to be **hereditary** if **whenever $a \in \Delta$, we have $a^\downarrow \subseteq \Delta$** , where $a^\downarrow = \{b \in \mathbb{A} \mid b \leq a\}$.
- The search for a set of distributions and weights for application of the Kuhn-Tucker criterion is decilitated by the fact that **any order consistent distribution is a convex combination of uniform distributions supported on connected hereditary subsets of \mathbb{A}** . Using convexity of the Kullback-Leibler divergence it is then realized that **we need only look at convex combinations of such uniform distributions**.

OVERVIEW OF CURRENT RESULTS

Trees of height 1. For a tree of height 1 as below,



the minimax predictor is given by $P^*(a_1) = \dots = P^*(a_{n-1})$ for all n , and

$$P^*(a_0) = \begin{cases} \frac{1}{k} & \text{when } a_k \leq n \leq b_k \\ 1 - \frac{n-1}{n-1+(1+\frac{1}{k-1})^{(k-1)k}} & \text{when } b_{k-1} \leq n \leq a_k \end{cases}$$

for certain sequences (a_k) and (b_k) satisfying $1 = b_1 < a_2 < b_2 < a_3 < b_3 < \dots$.

Matrix orderings. Suppose that $\mathbb{A} = \{a_{ij} \mid i, j = 1, \dots, n\}$. The ordering given by $a_{ii} \leq a_{ij}, a_{ji}$ for all i, j and no other relations is a (2-dimensional) matrix ordering. Again there is an exact result determining the minimax predictor, and again there is a peculiar variation with n .

Co-trees. Suppose that the ordering on \mathbb{A} defines a (suspended) co-tree, i.e. that there is a unique maximal element and that each element has at most 1 immediate successor. It is then possible to determine the minimax predictor by an algorithm which terminates in polynomial time (in the cardinality of \mathbb{A}).

COMMENTS ON OPEN PROBLEMS

More specific results for order structures. Given the peculiar variation with n of the minimax predictors for tree / matrix models, it seems unlikely that there is a general algorithm to cover all order structures. At the very least, it indicates that this would be rather complicated. However, **There are certainly more specific order models for which the monimax predictor can be determined, either in closed form or by a terminating algorithm.**

Relation of order models to other models. The matrix and tree orderings are interesting not only in their own right, but also because they are connected to **Bernoulli sources**. It turns out that the matrix ordering above have the same minimax predictor as the problem of predicting the first 2 characters from an n -letter Bernoulli source with no previous knowledge, and that trees of height 1 describe the problem of predicting 1 character, already knowing what came immediately before it. **Can this be extended? Are there other similar cases?**

General structural results. The remark concerning the calculation of the minimax predictor for \mathcal{P}_2 on the 'SIMPLE EXAMPLES' sheet purely from the knowledge of the minimax predictor for \mathcal{P}_1 can be extended and generalized. **Lacking an algorithm for general order structures, such general tricks could prove very effective, or at least an amusing curiosity.**